

# Forecasting intermittent data with complex patterns

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and Forecasting



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# Introduction

In the previous episodes (ISFs)...



(Pictures are from 'Rick and Morty' by Roiland and Harmon, 2013–2017)

# Introduction

Svetunkov and Boylan (2017) proposed an intermittent multiplicative state-space model.

We showed that this model underlies Croston (1972) and Teunter et al. (2011) methods.

We extended that model, presenting at ISF2017 the idea of using  $ETS(M,N,N)$ ,  $ETS(M,M,N)$  or  $ETS(M,Md,N)$  for demand sizes.



# Introduction

We showed how to select between different ETS models in this context.

The approach worked well on a WF wholesale data from Johnston et al. (1999).

The conclusion was: you can use one model for both intermittent and non-intermittent data.



# Motivation

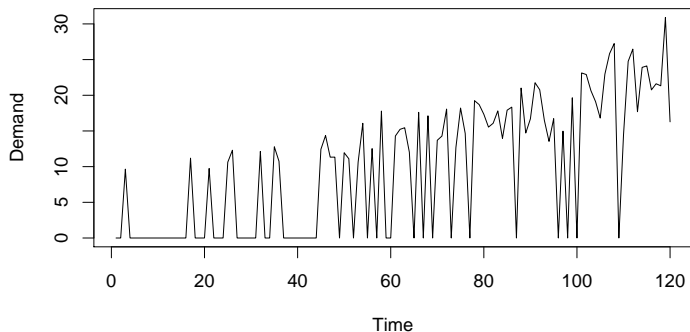
Now we want more!



# Motivation

The reason is this:

**Real time series example**



# Motivation

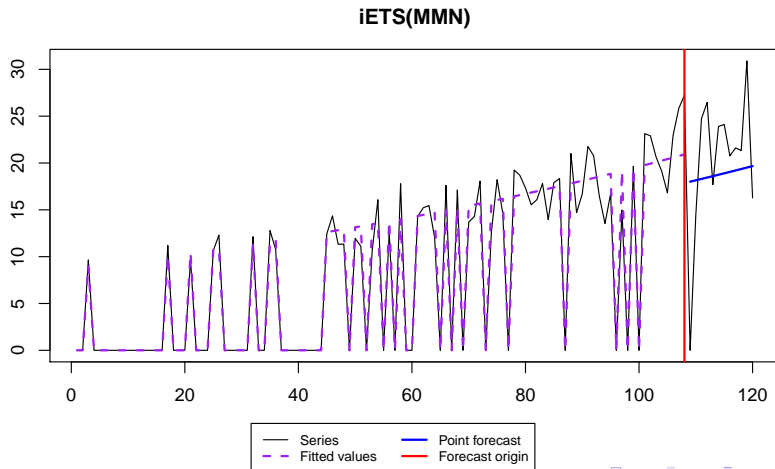
How do you deal with this type of data?

There is a trend, but the demand is intermittent.

We can predict the increase in demand sizes with  $iETS(M,M,N)_p...$



# Motivation

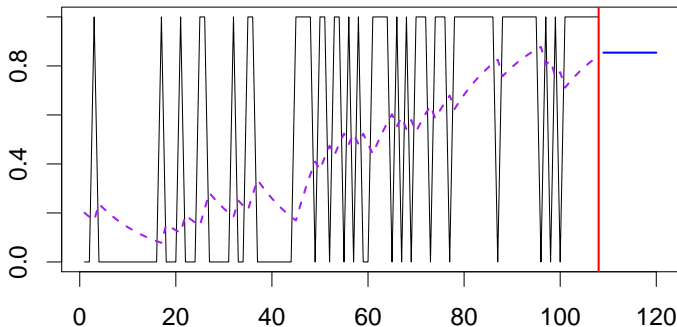




# Motivation

...Which underforecasts. Because we deal with the following:

## iSS, Probability-based

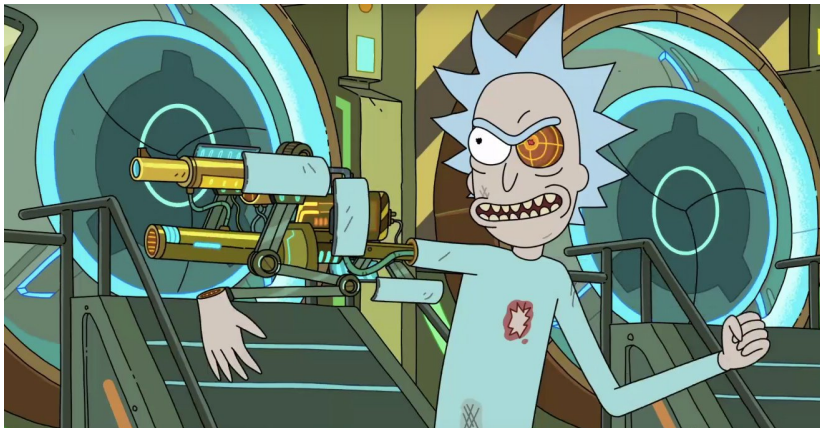


# Motivation

We need to capture complex patterns in the occurrence part of demand...



# The model



# iETS model

$$y_t = o_t z_t, \quad (1)$$

where  $o_t \sim \text{Bernoulli}(p_t)$ ,

$z_t$  is a **statistical model** of our choice

and  $p_t$  is another statistical model.



# Intermittent state-space model

Example.  $iETS(M,N,N)_p$  (with probability-based occurrence):

$$\begin{aligned}y_t &= o_t z_t \\z_t &= l_{t-1}(1 + \epsilon_t) \\l_t &= l_{t-1}(1 + \alpha \epsilon_t) \\o_t &\sim \text{Bernoulli}(p_t) \\p_t &= l_{p,t-1}(1 + \epsilon_{p,t}) \\l_{p,t} &= l_{p,t-1}(1 + \alpha_p \epsilon_{p,t})\end{aligned}\tag{2}$$



## Intermittent state-space model

Example.  $iETS(M,N,N)_p$  (with probability-based occurrence):

$$\begin{aligned}
 &y_t = o_t z_t \\
 &\left. \begin{aligned} z_t &= l_{t-1}(1 + \epsilon_t) \\ l_t &= l_{t-1}(1 + \alpha \epsilon_t) \end{aligned} \right\} \text{Demand sizes} \\
 &o_t \sim \text{Bernoulli}(p_t) \\
 &\left. \begin{aligned} p_t &= l_{p,t-1}(1 + \epsilon_{p,t}) \\ l_{p,t} &= l_{p,t-1}(1 + \alpha_p \epsilon_{p,t}) \end{aligned} \right\} \text{Demand occurrence}
 \end{aligned} \tag{3}$$

$1 + \epsilon_t \sim \log N(0, \sigma^2)$ , which means that  $z_t$  is always positive.

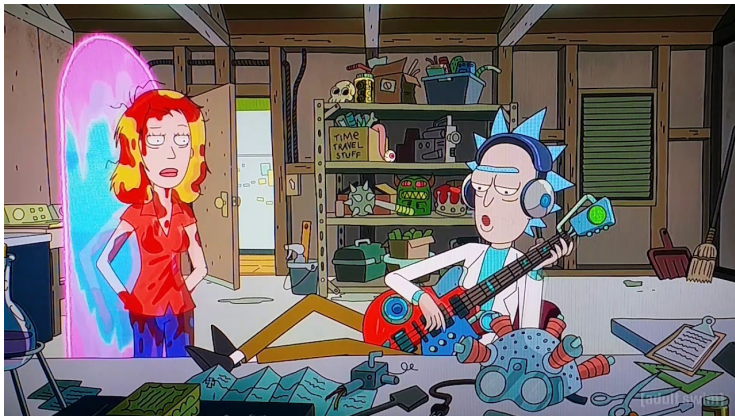
$1 + \epsilon_{p,t} \sim \log N(0, \sigma_p^2)$ .

So far, so good?



# The problem

But there is a tiny problem...



# The problem

The problems appear, when we start introducing additional components and variables in the model of  $p_t$ :

$$\begin{aligned} p_t &= l_{p,t-1} b_{p,t-1} (1 + \epsilon_{p,t}) \\ l_{p,t} &= l_{p,t-1} b_{p,t-1} (1 + \alpha_p \epsilon_{p,t}), \\ b_{p,t} &= b_{p,t-1} (1 + \beta_p \epsilon_{p,t}) \end{aligned} \tag{4}$$

$p_t$  should be in  $[0, 1]$

But if trend is positive,  $p_t$  might become greater than one.

Cutting off values is inhumane...





# Logistic transform

The solution - use a different model for  $p_t$ .

If we knew the true  $p_t$ , then we could use logit transform:

$$q_t = \log \left( \frac{p_t}{1 - p_t} \right) \quad (5)$$

$q_t$  is defined on  $(-\infty, \infty)$ .

We can use any model for  $q_t$ .



## Logistic transform

For example we can use ETS(A,A,N):

$$\begin{aligned}q_t &= l_{q,t-1} + b_{q,t-1} + \epsilon_{q,t} \\l_{q,t} &= l_{q,t-1} + b_{q,t-1} + \alpha_q \epsilon_{q,t}, \\b_{q,t} &= b_{q,t-1} + \beta_q \epsilon_{q,t}\end{aligned}\tag{6}$$

where  $\epsilon_{q,t} \sim \mathcal{N}(0, \sigma_q^2)$ .

We can extend this model with exogenous variables or seasonal components.

This would mean that in some cases the probability of occurrence increases / decreases.



# Logistic transform

In fact, we don't need to know either  $p_t$  or  $q_t$ , we only need to know  $\epsilon_{q,t}$ , initial values of  $l_{q,0}$  and  $b_{q,0}$ , and smoothing parameters values.

The latter four can be estimated... IF we have  $\epsilon_{q,t}$

But it is unobservable, so...



# Logistic model

One can only dream... right?



## Logistic transform. The rise of errors

There is a solution...

Use the inverse transform if the value  $\hat{q}_t$  is known:

$$\hat{p}_t = \frac{\exp(\hat{q}_t)}{1 + \exp(\hat{q}_t)} \quad (7)$$

Compare the predicted probability  $\hat{p}_t$  with the outcome  $o_t$ :

$$u_t = o_t - \hat{p}_t \quad (8)$$

The problem now is to translate this error into  $\epsilon_{q,t}$ .



## Logistic transform. The rise of errors

$u_t$  lies in  $(-1, 1)$ .

We transform  $u_t$ , so that it lies in  $(0, 1)$ :

$$u'_t = \frac{1+u_t}{2}.$$

and then use logit transform to obtain an estimate of error:

$$e_{q,t} = \log \left( \frac{1 + o_t - \hat{p}_t}{1 - o_t + \hat{p}_t} \right). \quad (9)$$

If  $o_t = \hat{p}_t$ , then  $e_{q,t} = 0$  (because  $o_t$  is binary).



# iETS<sub>l</sub> model

So the final model is:

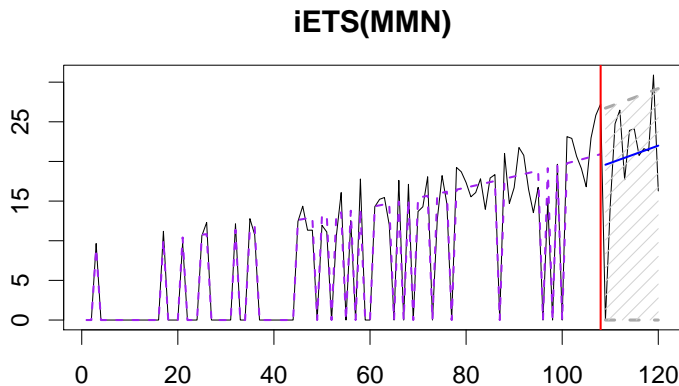
$$\begin{aligned}y_t &= o_t z_t \\z_t &\sim \text{ETS}(Y, Y, Y) \\o_t &\sim \text{Bernoulli}(p_t) \\p_t &= \frac{\exp(q_t)}{1 + \exp(q_t)}, \\q_t &\sim \text{ETS}(X, X, X) \\e_{q,t} &= \log\left(\frac{1 + o_t - \hat{p}_t}{1 - o_t + \hat{p}_t}\right)\end{aligned}, \quad (10)$$

where  $\text{ETS}(Y, Y, Y)$  is a multiplicative ETS, and  $\text{ETS}(X, X, X)$  is an additive one.



# iETS<sub>l</sub> model

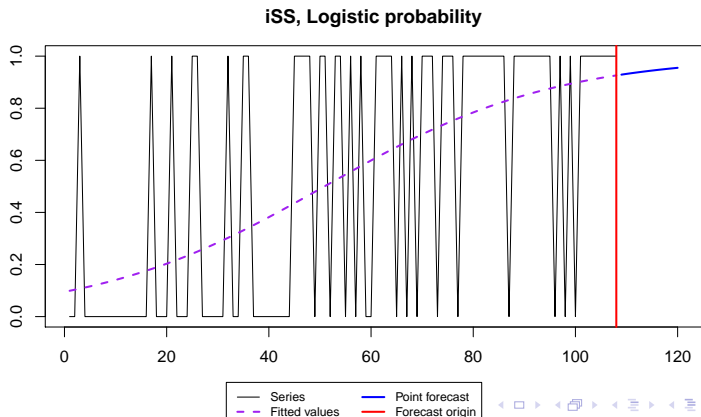
How does iETS<sub>l</sub> work?





# iETS<sub>I</sub> model

The occurrence part of iETS<sub>I</sub>



# iETS<sub>l</sub> model

So now we have:

- an extendable model...
- ...that captures complex patterns for demand sizes...
- ...AND demand occurrence part.



# Experiment

Yes! Now it is time for experiments...



# Experiment

Intermittent data of a Portuguese retailer.

5275 SKUs with at least 4 non-zero demands each.

Weekly data, 173 observations.

$h = \{1, 2, 3, 4\}$ .

Rolling with 52 origins.

sMSE, sME, sAPIS (Petropoulos and Kourentzes, 2015).



# Experiment

Croston, TSB and iMAPA from `tsintermittent` package

es function from `smooth` package with:

- $ETS(A,N,N)$ ,
- $iETS(M,Y,N)_f$  - fixed probability, selection of trend,
- $iETS(M,Y,N)_i$  - Croston style,
- $iETS(M,Y,N)_p$  - TSB style,
- $iETS(M,N,N)_l(A,N,N)$  - logistic with  $ETS(A,N,N)$  for occurrence,
- $iETS(M,Y,N)_l(A,N,N)$  - similar + selection of trend,
- $iETS(M,Y,N)_l(A,X,N)$  - similar + selection for occurrence.



# Results

Model	sME	sMSE	sAPIS
iETS(MYN) <sub>l</sub> (AXN)	0.460	<b>44.865</b>	<b>1.447</b>
iETS(MNN) <sub>l</sub> (ANN)	0.428	54.706	1.520
iETS(MYN) <sub>l</sub> (ANN)	0.471	54.744	1.512
iETS(MNN) <sub>p</sub>	0.596	55.391	1.493
ETS(AAN)	<b>0.176</b>	55.970	1.627
iETS(MNN) <sub>i</sub>	0.565	56.073	1.514
iMAPA	0.986	58.693	1.644
iETS(MNN) <sub>f</sub>	1.328	61.397	1.667
TSB	1.092	61.669	1.684
Croston	1.102	61.937	1.692



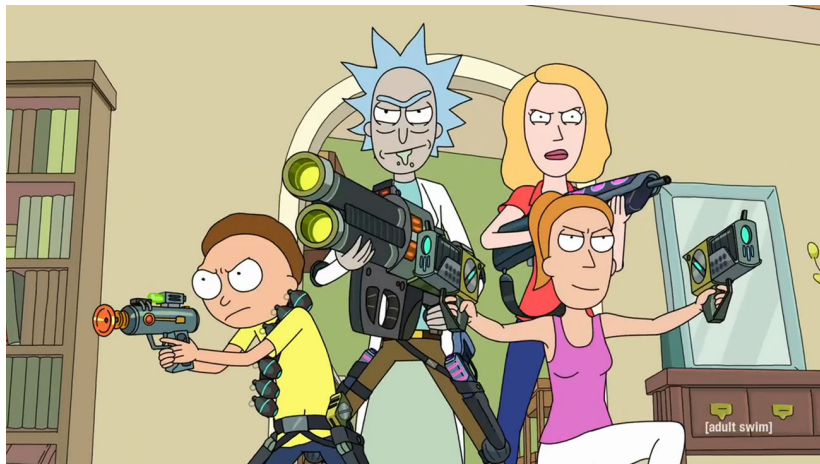
# Results

Based on the  $iETS_t$ :

Demand sizes	Demand occurrence		
	No trend	Trend	Overall
No trend	31.9%	35.7%	67.6%
Trend	19.0%	13.4%	32.4%
Overall	50.9%	49.1%	100%



# Conclusions





# Conclusions

- We now have a more general modelling framework;
- We have a new type of model for intermittent data;
- We can capture complex patterns in intermittent data;
- The approach seems to work well in practice.



## What's next?

- More thorough analysis of results driven by the data;
- Go multivariate – vector intermittent models:
  1. Extend the  $iETS_t$  to  $iVES_t$ ;
  2. Group time series based on the characteristics;
  3. Capture seasonality across similar time series;
  4. Introduce exogenous variables;
  5. Forecast groups of time series.
- In the end we should have a universal time series forecasting approach...



And that's how it's done!



# Thank you for your attention!

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